

Maximin-Projection Learning for Optimal Treatment Decision with Heterogeneous Individualized Treatment Effects

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Introduction

There are G groups of patients with

$$Y_g = h_g(X_g) + A_g(X_g^T \beta_g + c_0) + e_g.$$

Subgroup heterogeneity:

- ◆ Baseline function $h_g(\cdot)$.
- ◆ Propensity score function $\pi_g(\cdot)$.
- ◆ Individualized treatment effects β_g .
- ◆ Distribution of the error e_g .

Groupwise optimal treatment decision:

$$\mathbb{I}(x^T \beta_g + c_0 > 0).$$

Objective: Find a single treatment decision rule works reliably for future patients.

Motivation

Health assessment questionnaire progression data:

- ◆ Methotrexate combinations ($A_g = 1$) v.s methotrexate monotherapy ($A_g = 0$)
- ◆ Group 1: patients enrolled from 1990 to 1992
- ◆ Group 2: patients enrolled from 1993 to 1996
- ◆ Group 3: patients enrolled from 1997 to 2000
- ◆ **Heterogeneity due to patients' enrollment time**

	Group 1	Group 2	Group 3
$\beta_g^{(1)}$	0.05(0.11)	-0.40(0.17)	0.07(0.21)
$\beta_g^{(2)}$	0.07(0.11)	0.06(0.21)	0.32(0.16)

The schizophrenia data:

- ◆ Cognitive-behavioural therapy ($A_g = 1$) v.s supportive counselling ($A_g = 0$)
- ◆ Group 1: patients from Manchester
- ◆ Group 2: patients from Liverpool
- ◆ Group 3: patients from North Nottinghamshire
- ◆ **Heterogeneity due to patients' treatment centre**

	Group 1	Group 2	Group 3
$\beta_g^{(1)}$	1.35(10.21)	1.17(11.52)	-20.73(13.27)
$\beta_g^{(2)}$	7.87(10.39)	-10.56(8.84)	3.45(9.14)

Methods

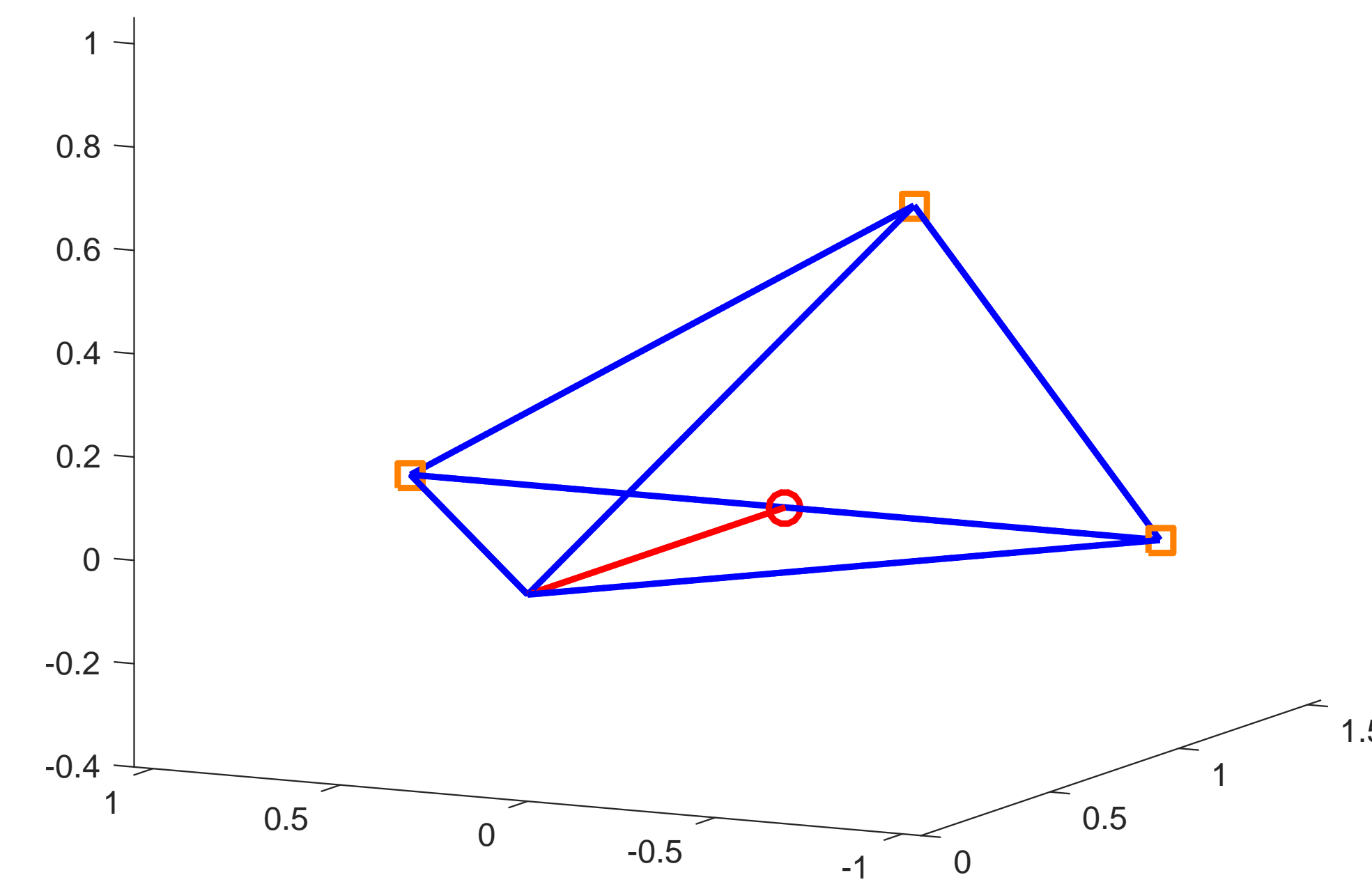
- ◆ **Random effects meta-analyses** (DerSimonian and Laird, 1986) estimate

$$\beta^R = \frac{1}{G} \sum_{g=1}^G \beta_g.$$

- ◆ **Consider the maximin effects:**

$$\beta^M = \arg \max_{\|\beta\|_2 \leq 1} \min_{g \in \{1, \dots, G\}} \beta^T \beta_g.$$

Geometrically, β^M corresponds to the **optimal equicorrelated point** of a subset of $\{\beta_1, \dots, \beta_G\}$.



- ◆ **Maximin-projection treatment decision:**

$$\mathbb{I}(x^T \beta^M + c^M > 0),$$

where $c^M = c_0 / \max_{\|\beta\|_2 \leq 1} \min_g \beta^T \beta_g$.

Statistical interpretation

Two optimality criteria:

- ◆ Value difference function:

$$VD_g(\beta, c) = E\{X_g^T \beta_g + c_0\} \mathbb{I}(X_g^T \beta > -c).$$
- ◆ Percentage of making correct decisions:

$$PCD_g(\beta, c) = 1 - E\{|\mathbb{I}(X_g^T \beta > -c) - d_g^{opt}(X_g)|\},$$
 where $d_g^{opt}(x) = \mathbb{I}(x^T \beta_g > -c_0)$.

Theorem 1: Under certain conditions,

$$(\beta^M, c^M) = \arg \max_{\beta, c} \min_{g \in \{1, \dots, G\}} VD_g(\beta, c),$$

$$(\beta^M, c^M) = \arg \max_{\beta, c} \min_{g \in \{1, \dots, G\}} PCD_g(\beta, c).$$

- ◆ **Observed datasets:**

$$\{(X_{gj}, A_{gj}, Y_{gj}), g = 1, \dots, G, j = 1, \dots, m_g\}.$$

- ◆ **Parameter estimation:**

- ◆ **Step 1:** Posit some parametric model $\pi_g(\cdot, \alpha_g)$ for the $\pi_g(\cdot)$ and $h_g(\cdot, \eta_g)$ for $h_g(\cdot)$. Estimate parameters by jointly solving:

$$\sum_{j=1}^{m_g} \frac{\partial \pi_g(X_{gj}, \alpha_g)}{\partial \alpha_g} \frac{\{A_{gj} - \pi_g(X_{gj}, \alpha_g)\}}{\pi_g(X_{gj}, \alpha_g) \{1 - \pi_g(X_{gj}, \alpha_g)\}} = 0,$$

$$\sum_{j=1}^{m_g} \frac{\partial h_g(X_{gj}, \eta_g)}{\partial \eta_g} \varepsilon_{gj}(\eta_g, \beta_g, c_0) = 0,$$

$$\sum_{j=1}^{m_g} X_{gj} \{A_{gj} - \pi_g(X_{gj}, \alpha_g)\} \varepsilon_{gj}(\eta_g, \beta_g, c_0) = 0,$$

$$\sum_{j=1}^{m_g} \sum_{g=1}^G \{A_{gj} - \pi_g(X_{gj}, \alpha_g)\} \varepsilon_{gj}(\eta_g, \beta_g, c_0) = 0.$$

where

$$\varepsilon_{gj}(\eta_g, \beta_g, c_0) = \{Y_{gj} - h_g(X_{gj}, \eta_g) - A_{gj}(X_{gj}^T \beta_g + c_0)\}.$$

- ◆ **Step 2:** Estimate β^M and c^M by

$$\hat{\beta}^M = \arg \max_{\|\beta\|_2 \leq 1} \min_{g \in \{1, \dots, G\}} \beta^T \hat{\beta}_g,$$

$$\hat{c}^M = \hat{c}_0 / \min_g \hat{\beta}_g^T \hat{\beta}_g^M.$$

Step 2 is a QCLP (Lee et al., 2016) and can be efficiently computed. Output

$$d_M(x) = \mathbb{I}(x^T \hat{\beta}^M + \hat{c}^M > 0).$$

- ◆ **Step 3:** Inference for β^M and c^M via bootstrap.

- 1 Independently generate B bootstrap samples within each group,

$$\{(X_{gj}^{(b)}, A_{gj}^{(b)}, Y_{gj}^{(b)}), g = 1, \dots, G, j = 1, \dots, m_g\}, \quad (1)$$

for $b = 1, \dots, B$.

- 2 Calculate $\hat{\beta}^{M(b)}$ and $\hat{c}^{M(b)}$ based on (1).
- 3 Obtain confidence intervals based on quantiles of $(\hat{\beta}^{M(1)}, \dots, \hat{\beta}^{M(B)})$ and $(\hat{c}^{M(1)}, \dots, \hat{c}^{M(B)})$.

Statistical property

Theorem 2: Under certain conditions, we have

$$\hat{\beta}^M \xrightarrow{P} \beta^M \text{ and } \hat{c}^M \xrightarrow{P} c^M.$$

In addition, $\sqrt{n}(\hat{\beta}^M - \beta^M)$ and $\sqrt{n}(\hat{c}^M - c^M)$ are jointly asymptotically normal with mean zero and some covariance matrix V^M .

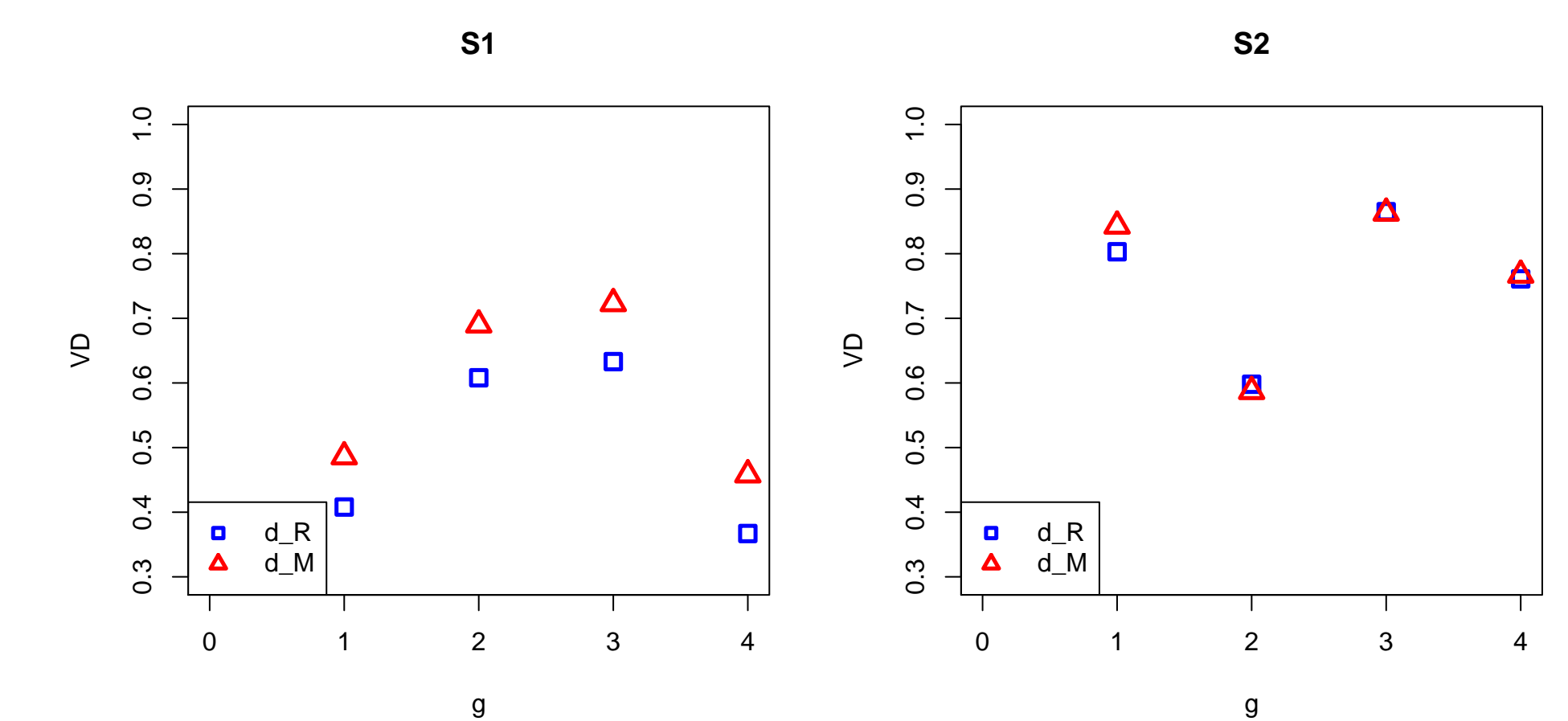
Simulations

- ◆ **Two scenarios** with different β_g 's, $G = 4$.

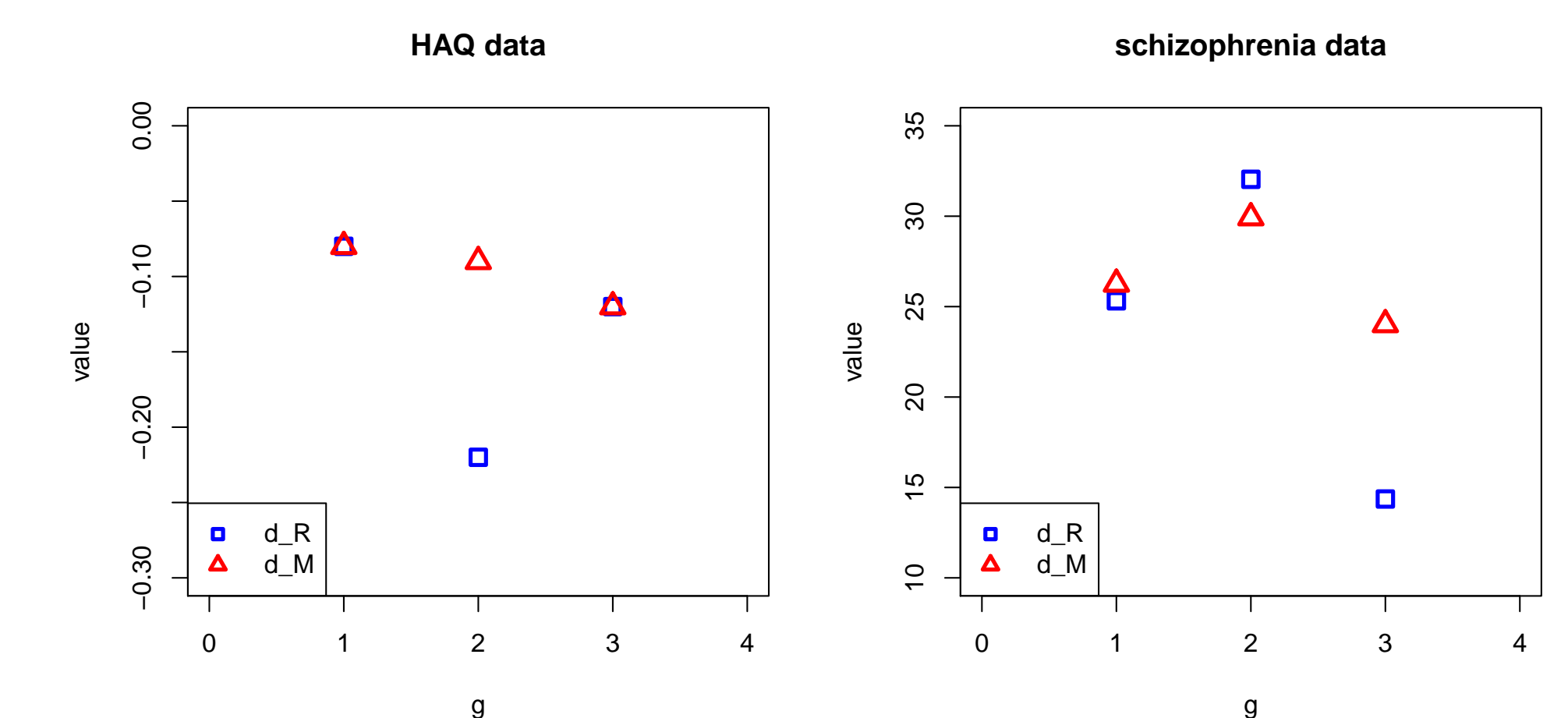
- ◆ **Scenario 1:** $\beta_1 = (2, 0)^T$, $\beta_2 = 2(\cos(15^\circ), \sin(15^\circ))^T$, $\beta_3 = 2(\cos(70^\circ), \sin(70^\circ))^T$, $\beta_4 = (0, 2)^T$.

- ◆ **Scenario 2:** $\beta_1 = 2.2(\cos(30^\circ), \sin(30^\circ))^T$, $\beta_2 = 1.5(\cos(45^\circ), \sin(45^\circ))^T$, $\beta_3 = 2.2(\cos(54^\circ), \sin(54^\circ))^T$, $\beta_4 = 2(\cos(60^\circ), \sin(60^\circ))^T$.

- ◆ Compare $d_M(\cdot)$ with the treatment regime based on random effects meta-analyses $d_R(\cdot)$.



Real data application



Contributions

- ◆ Propose a maximin-projection learning to aggregate optimal treatment decisions for patients from different populations with heterogeneity.
- ◆ Show the proposed treatment decision has nice statistical interpretation in the sense of maximizing the minimum PCD and value difference function.
- ◆ Provide a geometrical characterization of the maximin estimator.
- ◆ Study the statistical properties of the estimators.